

Heavy-tailed density regression using the blended generalised Pareto distribution and xSPQR

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Extrapolation — peaks-over-threshold

- **Pickands–de Haan–Balkema Theorem:**

High threshold excesses $Y - u \mid Y > u$ may be approximated by the generalised Pareto (GP) distribution: if there exists a scaling function $a(u) > 0$ such that as $u \rightarrow y_F$ (upper endpoint)

$$\Pr \left(\frac{Y - u}{a(u)} > y \mid Y > u \right) \rightarrow \text{a non-degenerate distribution}$$

$$\text{then it must be } 1 - F_{\text{GP}}(y \mid \sigma_u, \xi) := \begin{cases} (1 + \xi y / \sigma_u)_+^{-1/\xi}, & \xi \neq 0, \\ \exp(-y / \sigma_u), & \xi = 0, \end{cases},$$

where $\sigma_u > 0$ and $\xi \in \mathbb{R}$.

- In practice, we model excesses directly as $(Y - u) \mid Y > u \sim \text{GP}(\sigma_u, \xi)$ where u is some high pre-specified threshold.

Regression

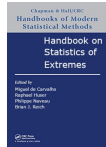
What if we have covariates $\mathbf{X} \in \mathbb{R}^p$?

- Often make parametric assumptions about $Y|\mathbf{X} = \mathbf{x}$, e.g.,

$$(Y - u(\mathbf{x})) | (\mathbf{X} = \mathbf{x}, Y > u(\mathbf{x})) \sim \text{GP}(\sigma_u(\mathbf{x}), \xi(\mathbf{x})),$$

with $u(\mathbf{x}) > 0$ some varying threshold function.

- Lots of AI-based options:**
 - neural networks**, e.g., Allouche et al. (2024), Cisneros et al. (2024), Pasche and Engelke (2024), Richards and Huser (2025).
 - trees (Farkas et al., 2024) and forests (Gnecco et al., 2024)
 - boosting (Velthoen et al., 2023; Koh, 2023)
 - GAMs (Chavez-Demoulin and Davison, 2005; Youngman, 2019)
- Richards, J. and Huser, R. (2025). **Extreme Quantile Regression with Deep Learning**. In *Handbook on Statistics of Extremes*, Chapman & Hall/CRC.



Conditional setting

What if we have covariates $\mathbf{X} \in \mathbb{R}^p$?

- Often make parametric assumptions about $Y|\mathbf{X} = \mathbf{x}$, e.g.,

$$(Y - u(\mathbf{x})) \mid (\mathbf{X} = \mathbf{x}, Y > u(\mathbf{x})) \sim \text{GP}(\sigma_u(\mathbf{x}), \xi(\mathbf{x})),$$

with $u(\mathbf{x}) > 0$ some varying threshold function.

- What about i) below the threshold, ii) choosing the threshold, iii) interpretability?
- We propose a semi-parametric density regression model that has GP upper-tails **without the need for threshold selection**.

Background – SPQR

Introduced by Xu and Reich (2021), SPQR is a flexible, semi-parametric approach to **conditional density estimation**.

- **No parametric assumptions**; instead, the conditional density is a convex combination of M -spline basis functions:

$$f_{\text{SPQR}}(y|\mathbf{x}) = \sum_{k=1}^K w_k(\mathbf{x}) M_k(y),$$

with weights $w_k(\mathbf{x}) : \mathbb{R}^p \mapsto [0, 1]$, $k = 1, \dots, K$, satisfying $\sum_{k=1}^K w_k(\mathbf{x}) = 1$ for all $\mathbf{x}$.

- Each basis function, $M_k(y)$, is a **valid PDF** on $[0, 1]$ (Ramsay, 1988).
- The integral of an M -spline is an I -spline:

$$F_{\text{SPQR}}(y|\mathbf{x}) = \sum_{k=1}^K w_k(\mathbf{x}) I_k(\mathbf{x}).$$

- The weights $\mathcal{W}(\mathbf{x}) := \{w_1(\mathbf{x}), \dots, w_K(\mathbf{x})\}$ are modelled as a MLP with softmax final layer.
- Although very flexible, and fast-to-compute, F_{SPQR} satisfies **no asymptotic guarantees and has bounded support**.

Blended Generalised Pareto

- Castro-Camilo et al. (2022) proposed the **blended generalised extreme value distribution (bGEV)**, which blends the **Gumbel** and **Fréchet** distributions
 - ⇒ the resulting distribution function has an exact **Gumbel** lower-tail and **Fréchet** upper-tail.
- We follow a similar idea, but instead **blend the GP distribution with a constituent *bulk* distribution**, say F_{bulk} .
- Here we present the specific case of the **unconditional** blended GP with $F_{\text{bulk}} := F_{\text{SPQR}}$; we will introduce covariates later.



Play along at home!

You can also follow the link
[https://reetamm-xspqr.share.
connect.posit.cloud](https://reetamm-xspqr.share.connect.posit.cloud)

Blended Generalised Pareto

We define a bGP($\mathcal{W}, \xi$) r.v. via its **continuous** distribution function

$$H(y|\mathcal{W}, \xi) = \begin{cases} F_{\text{SPQR}}(y|\mathcal{W})^{1-p(y)} F_{\text{GP}}(y - \tilde{u}|\tilde{\sigma}_u, \xi)^{p(y)}, & y > \tilde{u}, \\ F_{\text{SPQR}}(y|\mathcal{W}), & y \leq \tilde{u}, \end{cases} \quad (1)$$

where $p(y) \in [0, 1]$ is a weighting function;

$$p(y) = p(y; a, b, c_1, c_2) = F_{\text{Beta}}\left(\frac{y-a}{b-a}, c_1, c_2\right),$$

where $F_{\text{Beta}}(\cdot, c_1, c_2)$ is a Beta(c_1, c_2) dist. with shapes $c_1 > 3, c_2 > 3$.

Note that $p(y) = 0$ for any $y < a$ and $p(y) = 1$ for any $y > b$.

Blended generalised Pareto

- We blend F_{SPQR} and F_{GP} in the interval $[a, b] \subset [0, 1]$, where the bounds are the p_a and p_b quantiles of F_{SPQR} ($p_b > p_a$).
- To ensure continuity of H , we require

$$\begin{aligned}p_a &:= F_{\text{SPQR}}(a|\mathcal{W}) = F_{\text{GP}}(a - \tilde{u}|\tilde{\sigma}_u, \xi) \\p_b &:= F_{\text{SPQR}}(b|\mathcal{W}) = F_{\text{GP}}(b - \tilde{u}|\tilde{\sigma}_u, \xi),\end{aligned}$$

with:

$$(\tilde{\sigma}, \tilde{u}) = \begin{cases} \left(\frac{\xi(a-b)}{(1-p_a)^{-\xi} - (1-p_b)^{-\xi}}, a - \frac{(a-b)\{(1-p_a)^{-\xi} - 1\}}{(1-p_a)^{-\xi} - (1-p_b)^{-\xi}} \right), & \xi \neq 0, \\ \left(\frac{(a-b)}{\log(1-p_a) - \log(1-p_b)}, a - \frac{(a-b)\{-\log(1-p_a)\}}{\log(1-p_a) - \log(1-p_b)} \right), & \xi = 0, \end{cases};$$

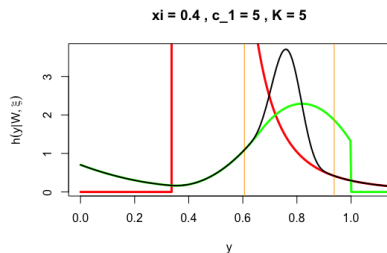
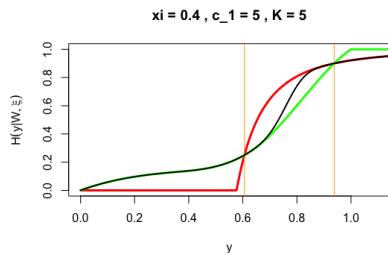
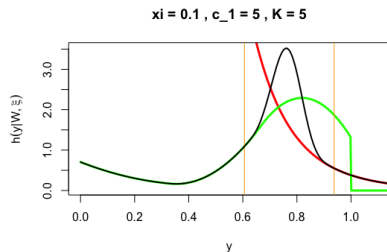
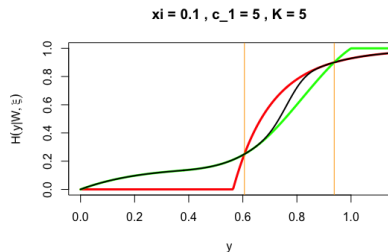
note that $\tilde{u} < a$.

Blended GP

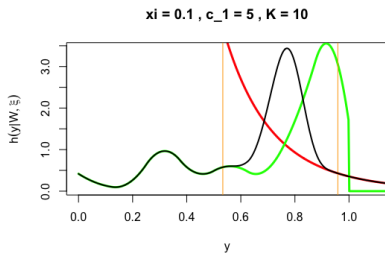
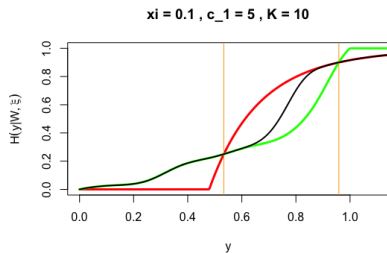
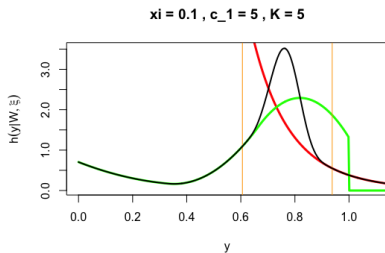
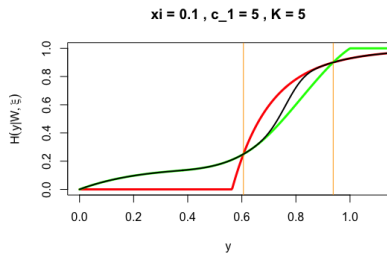
- For $\xi < 0$, the upper-endpoint of $H(\cdot|\mathcal{W}, \xi)$ satisfies $\tilde{u} - \tilde{\sigma}_u/\xi > b$;
for $\xi \geq 0$, the upper-endpoint of $H(\cdot|\mathcal{W}, \xi)$ is infinite.
- The density is **closed-form**, and is **smooth** and **continuous**.

Increasing tail-heaviness

Left: bGP, GP, SPQR distribution. Right: corresponding density functions.

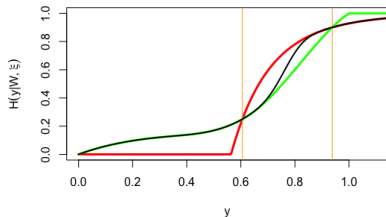


Increasing bulk-flexibility

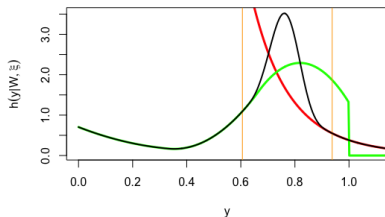


Increasing SPQR weighting

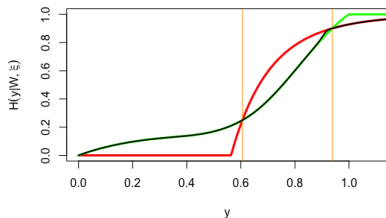
$\xi = 0.1, c_1 = 5, K = 5$



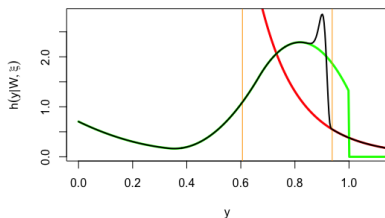
$\xi = 0.1, c_1 = 5, K = 5$



$\xi = 0.1, c_1 = 50, K = 5$

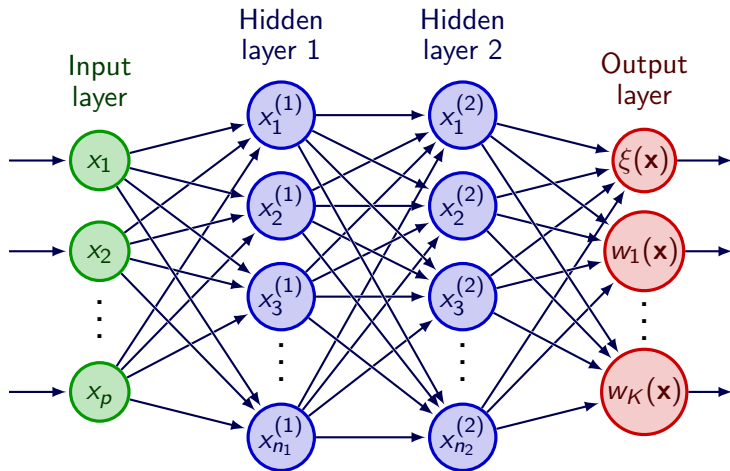


$\xi = 0.1, c_1 = 50, K = 5$



xSPQR

In the presence of covariates, we model $\mathbf{x} \mapsto (\xi(\mathbf{x}), \mathcal{W}(\mathbf{x}))$ via an MLP:



We refer to this framework as extremal SPQR (xSPQR).

Inference/covariate importance

- Inference proceeds via maximum likelihood using Adam.
- xSPQR can be pre-trained with an SPQR fit.
- Via the R interface to keras.
- Variable importance (VI) can be assessed for conditional quantile function $Q(\tau|\mathbf{x})$ at $\tau \in (0, 1)$ **separately** of the shape $\xi(\mathbf{x})$
- Using model-agnostic accumulated local effects (ALEs; Apley and Zhu, 2020).

Simulation study

- Covariates $\mathbf{X}_i, i = 1, \dots, 3$, are independent $\text{Unif}(0,1)$.
- Response $Y \mid (\mathbf{X} = \mathbf{x})$ is $\text{log-normal}(\mu(\mathbf{x}), \sigma(\mathbf{x}))$ with

$$\mu(\mathbf{x}) = 5(1 - 1/[1 + \exp\{-(1 - 5x_1x_2)\}])$$

and

$$\sigma(\mathbf{x}) = 1/[1 + \exp\{-(1 - 5x_1x_2)\}].$$

- Only X_1 and X_2 act on Y .
- We take the MLP to have two layers, with n_h nodes and sigmoid activation in each layer.

Simulation study

- To evaluate estimation accuracy, we compute the **integrated conditional 1-Wasserstein distance (IWD)**

$$\text{IWD} = \int_{\mathcal{X}} \int_0^1 |Q(y|\mathbf{x}) - \hat{Q}(y|\mathbf{x})| d\mathbf{x},$$

where $\mathcal{X}$ is the sample space for $\mathbf{X}$ and $Q(y|\mathbf{x})$ denotes the conditional quantile function.

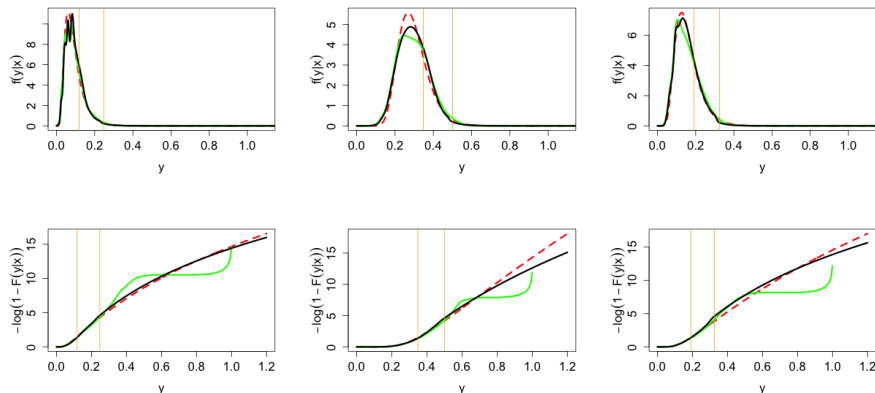
- We also consider a **tail-calibrated version** of the IWD, denoted by **tlIWD**, which is constructed by replacing the limits of the inner integral of (17) with $[0.999, 1]$.

Results

n	K	n_h	tIWD	(p_a, p_b, c_1)
1000	15	16	11.2 (10.3, 12.3)/ 9.23 (7.99, 10.6)	(0.9, 0.999, 5)
	15	32	9.50 (8.63, 10.6)/ 9.66 (8.18, 11.2)	(0.925, 0.999, 5)
	25	16	12.0 (11.2, 13.0)/ 9.56 (8.42, 10.9)	(0.925, 0.999, 5)
	25	32	9.20 (8.31, 9.96)/ 9.80 (8.70, 11.1)	(0.925, 0.999, 5)
10000	15	16	10.6 (9.51, 11.3)/ 7.08 (6.40, 7.85)	(0.75, 0.99, 25)
	15	32	10.7 (10.0, 11.6)/ 6.99 (6.36, 8.05)	(0.75, 0.99, 25)
	25	16	8.60 (7.33, 10.0)/ 5.56 (4.45, 6.86)	(0.75, 0.99, 25)
	25	32	10.2 (9.40, 16.6)/ 5.29 (4.59, 6.50)	(0.75, 0.99, 25)

Median (25%, 75% quantiles) of tIWD estimates are reported for the original/**heavy-tailed** SPQR model, with the hyper-parameters (p_a, p_b, c_1) optimised for each row. Lower values are better.

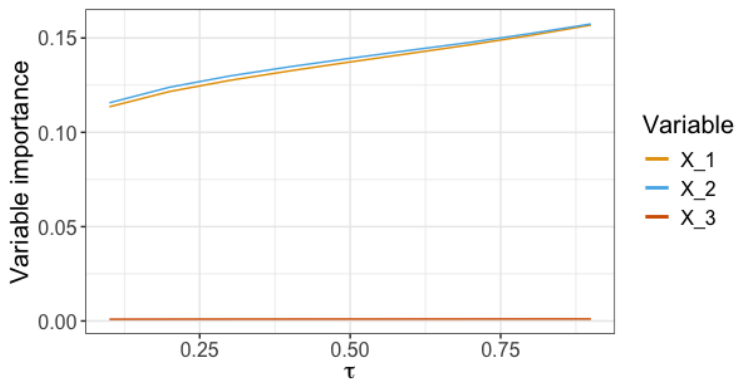
Test density estimation



Density (top) and log-survival (bottom) functions.

True, **SPQR**, and **xSPQR**.

Variable importance



VI scores (10^{-2}) for $\xi(\mathbf{x})$: 1.56, 2.34, 0.09.

Case study: US wildfire burnt areas

- Burnt areas for over 10,000 moderate and large wildfires in the US, 1990–2020 (Lawler and Shaby, 2024).
- First and last 5 years used for testing. Model trained for 1995–2015.
- This leaves **6416 fires** for training and **3344 fires** for testing.

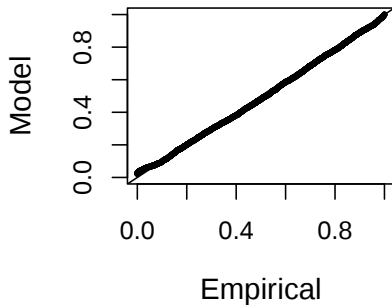
Lawler, E. S. and Shaby, B. A. (2024). Anthropogenic and meteorological effects on the counts and sizes of moderate and extreme wildfires. *Environmetrics*, 35(7):e2873

Case study: US wildfire burnt areas

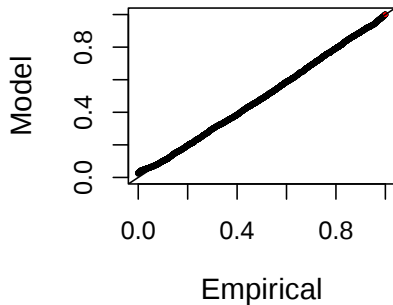
- We model the impacts of $\mathbf{X} =$
 - `pr_prev`: total precip. last year;
 - `pr_curr`: total precip. this month;
 - `rmin`: relative humidity;
 - `tmax`: maximum temperature;
 - `wspd`: windspeed;
 - `fire_yr`: fire year;on $Y = \sqrt{\text{Burnt area}}$.
- Model hyper-parameters/MLP architecture optimised via grid-search:
 - We here use a 2-layered MLP with $N_h = 12$ nodes per layer, sigmoid activations, and $K = 25$ basis functions.
 - We also constrain $\xi(\mathbf{x}) > 0$.

Model fits - bulk

SPQR

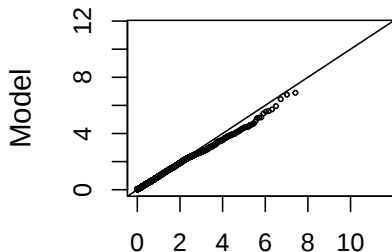


xSPQR



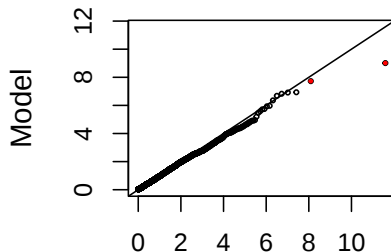
Model fits - tail

SPQR



Empirical

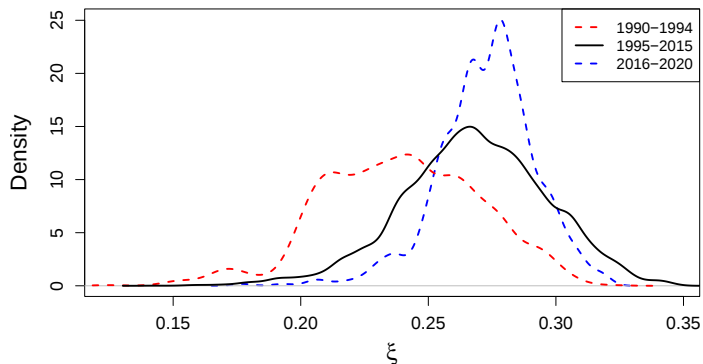
xSPQR



Empirical

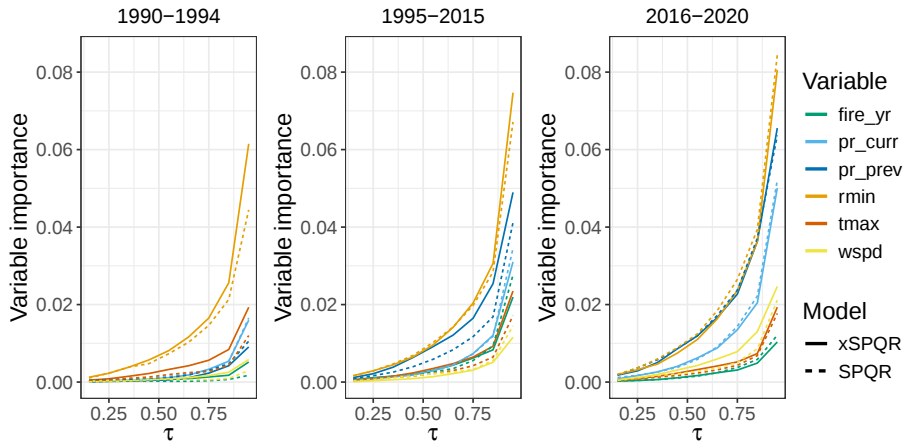
Red test points are impossible with SPQR.

Estimates of $\xi(\mathbf{x})$



Density of the estimated $\xi(\mathbf{x})$, stratified by time period.

Relative variable importance - bulk



Relative variable importance - tail

Time period	pr_prev	rmin	tmax	wspd	pr_curr	fire_yr
1990–1994	2.35	2.08	1.41	0.90	0.91	0.22
1995–2015	2.87	2.13	1.82	1.72	1.21	0.61
2016–2020	2.39	1.69	1.42	2.26	1.58	0.50

Spatial variation in quantiles

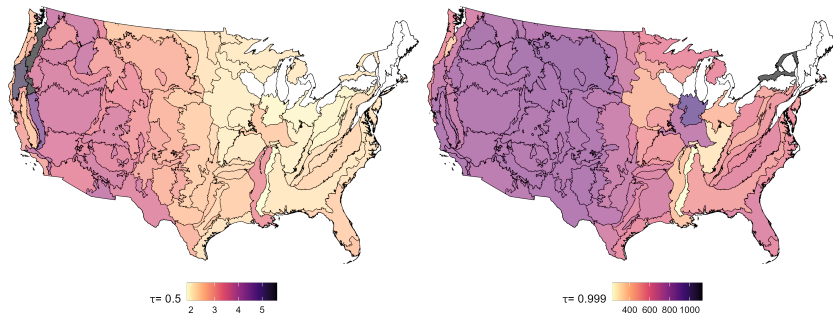


Figure: Estimates of the median (left) and 0.999-quantile (right) of burnt area (in 1000s of acres) for all observed wildfires, averaged over L3 ecoregions. Transparent regions do not include any observed wildfires.

Conclusion

- Very flexible density regression model that is **EVT-compliant**.
- Requires no modelling of an intermediate exceedance threshold and provides a characterisation of the full density.
- Easily extendable to full real support and lower-tailed GP.
- Majumder, R. and Richards, J. (2025+). Semi-parametric bulk and tail regression using spline-based neural networks. [arxiv:2504.19994](https://arxiv.org/abs/2504.19994).



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Bridging Heavy Tails and Artificial Intelligence

 Participating journal: [Extremes](#)

 Closed for submissions

The increasing frequency and severity of extreme events—such as catastrophic flooding, record-breaking temperatures, and unprecedented heatwaves—has highlighted the urgent need for innovative approaches in risk assessment and modeling. Modern advancements in data-collection techniques have provided increasingly large and complex datasets, which can only be processed using fast and scalable algorithms and computational software. This special issue aims to bridge the gap between Artificial Intelligence (AI) and Extreme Value Theory (EVT) to harness the strengths of both fields and address the growing challenges posed by these extreme events.

 Participating journal



Journal
Extremes
Extremes is a dedicated platform for the publication of original research in statistical extreme value theory and its applications across various fields.

Publishing model	Hybrid
Journal Impact Factor	1.1 (2023)
Downloads	58k (2024)
Submission to first decision (median)	7 days

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Appendix: construction of M-splines

Defined on a set of $K + d$ knots, $t_1, \dots, t_{K+d}$, which we take to be empirical quantiles of the training Y with equally-spaced levels.

For $d = 1$,

$$M_k(y|d) = \begin{cases} \frac{1}{t_{k+1} - t_k}, & t_k \leq y < t_{k+1}, \\ 0, & \text{otherwise.} \end{cases}$$

and, for $d > 1$,

$$M_k(y|d) = \frac{d[(y - t_k)M_k(y|d-1) + (t_{k+d} - y)M_{k+1}(y|d-1)]}{(d-1)(t_{k+d} - t_k)}.$$

For SPQR/xSPQR, $d = 3$.

Appendix: variable importance scores

Consider a generic differentiable function $g(\mathbf{x})$, where $\mathbf{x} = (x_1, \dots, x_p)$ is the vector of covariates. The sensitivity of $g(\mathbf{x})$ to covariate x_j is quantified by the partial derivative

$$\dot{g}_j(x_j) = \frac{\partial g(\mathbf{x})}{\partial x_j}.$$

The accumulated local effect (ALE) of x_j on $g(\cdot)$ is then defined as

$$\text{ALE}_j(x_j; g) = \int_{z_{0,j}}^{x_j} \mathbb{E}[\dot{g}_j(x_j) | x_j = z_j] dz_j,$$

where $z_{0,j}$ is an approximate lower bound for x_j .

Appendix: variable importance scores (cont.)

Following Greenwell et al. (2018), we measure **heterogeneity of the effect of X_j on $g(\cdot)$** by taking the **standard deviation of $ALE_j(X_j; g)$ with respect to X_j** .

The variable importance (VI) score for X_j on $g(\cdot)$ is

$$VI_j(g) = \sqrt{\text{Var}_{X_j}[ALE_j(X_j; g)]}.$$

For xSPQR, replace $g(\cdot)$ with the conditional τ -quantile function or $\xi(\mathbf{x})$.