

A DEEP LEARNING APPROACH TO MODELLING JOINT ENVIRONMENTAL EXTREMES

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Motivation

- Quantifying probabilities of joint occurrence of extremes is important in risk analysis and applications in, e.g., finance, climate, neuroscience.
- Many models for multivariate extremal dependence make restrictive assumptions about joint tail decay and do not scale well to high dimensions.
- *Geometric extremes* [1] is a flexible framework for modelling multivariate extremes that considers the deterministic limiting shape of scaled sample clouds.
- Estimation in this framework is i) limited to a low-dimensional setting, ii) requires restrictive parametric assumptions, or iii) provides estimates that do not satisfy **theoretical properties**.
- We provide new results for *Geometric extremes* that allow us to construct estimators that produce valid model estimates. We exploit these in a deep learning framework to build **flexible** and **scalable** models for multivariate extremal dependence.

Geometric representation of multivariate extremes

- For a d -dimensional random vector $\mathbf{X} \in \mathbb{R}^d$ on standard Laplace margins, with density $f(\cdot)$.
- Given n independent realisations of \mathbf{X} , consider the **scaled sample cloud**

$$C_n := \{\mathbf{X}_i/r_n; i = 1, \dots, n\},$$

as $n \rightarrow \infty$, where r_n is a suitably chosen normalising sequence.

- Under mild conditions C_n converges in probability onto the star-shaped and compact set

$$\mathcal{G} := \{\mathbf{x} \in \mathbb{R}^d : g(\mathbf{x}) \leq 1\} \subseteq [-1, 1]^d,$$

which intersects the boundary of $[-1, 1]^d$ at least once in each component.

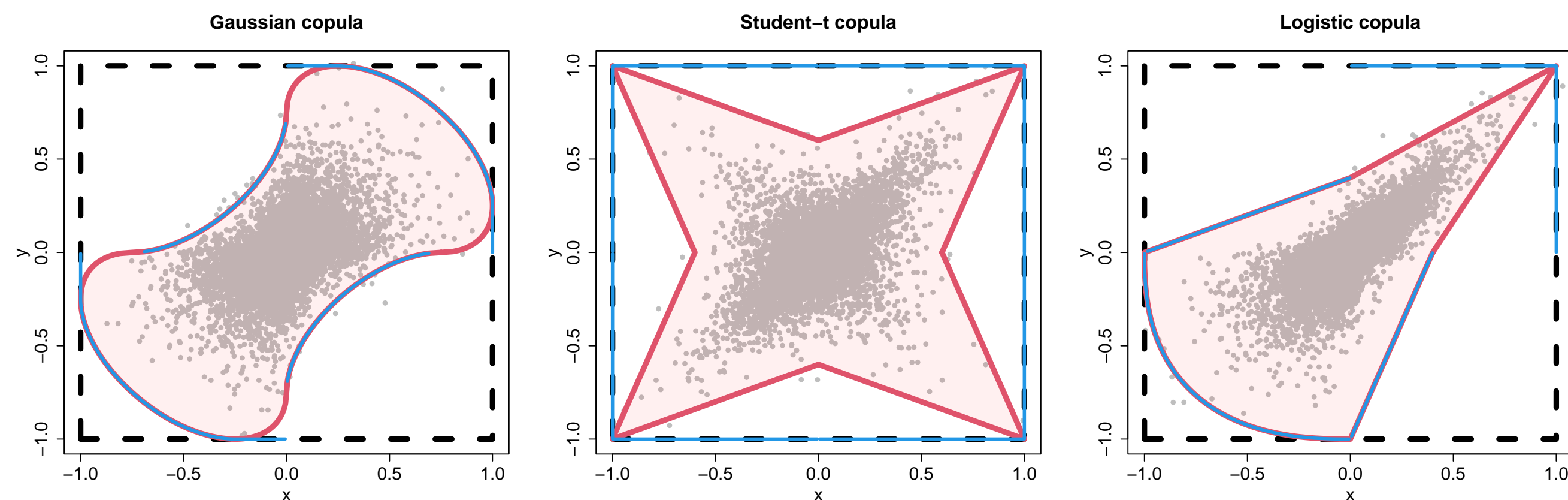


Figure 1. Scaled sample clouds of size $n = 10,000$. Shaded regions and solid red lines give the limit set \mathcal{G} and its boundary $\partial\mathcal{G}$, while the blue lines denote the set $\{\mathbf{w}/\Lambda(\mathbf{w}) : \mathbf{w} \in \mathcal{S}^{d-1} \setminus \mathcal{A}\}$.

- A sufficient condition for convergence onto \mathcal{G} [2] is that

$$-\log f(t\mathbf{x}) \sim tg(\mathbf{x}), \quad t \rightarrow \infty, \quad \mathbf{x} \in \mathbb{R}^d,$$

where $g(\cdot)$ is a continuous 1-homogeneous **gauge function** on \mathbb{R}^d .

- [3] and [1] link $\partial\mathcal{G}$ to popular extremal dependence models. We can also use $\partial\mathcal{G}$ to get probabilities.
- For example, [4] assume that, as $u \rightarrow \infty$,

$$\Pr\left(\min_{i=1, \dots, d} \{X_i\} > u\right) = L(e^u) \exp(-u/\eta),$$

where $L(\cdot)$ is slowly-varying and $\eta \in (0, 1]$ is the **coefficient of tail dependence**. Under asymptotic dependence, we have $\eta = 1$ and $\lim_{u \rightarrow \infty} L(u) > 0$.

- To connect this to $\partial\mathcal{G}$, we have that

$$\eta = \min \left\{ s \in (0, 1] : [s, \infty]^d \cap \partial\mathcal{G} = \emptyset \right\}.$$

- Define (R, \mathbf{W}) by $\mathbf{X} \mapsto (R, \mathbf{W}) := (\|\mathbf{X}\|, \mathbf{X}/\|\mathbf{X}\|)$ for $R > 0$ and $\mathbf{W} \in \mathcal{S}^{d-1}$, where $\mathcal{S}^{d-1} := \{\mathbf{x} \in \mathbb{R}^d : \|\mathbf{x}\| = 1\}$ denotes the unit $(d-1)$ -sphere and $\|\cdot\|$ is the Euclidean norm.

- Then

$$\partial\mathcal{G} := \{r\mathbf{w} : r > 0, \mathbf{w} \in \mathcal{S}^{d-1}, g(r\mathbf{w}) = 1\} = \{\mathbf{w}/g(\mathbf{w}) : \mathbf{w} \in \mathcal{S}^{d-1}\},$$

and so we only require a model for $g(\cdot)$ to estimate $\partial\mathcal{G}$.

Theoretical developments

- **Proposition:** For all $\mathbf{w} \in \mathcal{S}^{d-1}$, the gauge function $g(\cdot)$ satisfies the constraint that

$$g(\mathbf{w}) \geq \|\mathbf{w}\|_\infty,$$

where $\|\mathbf{x}\|_\infty := \max\{|x_1|, \dots, |x_d|\}$ denotes the infinity norm.

- Suppose we have **any** $h(\cdot) : \mathcal{S}^{d-1} \rightarrow \mathbb{R}_+$ that satisfies $1/h(\mathbf{w}) \geq \|\mathbf{w}\|_\infty$ for all $\mathbf{w} \in \mathcal{S}^{d-1}$.

Proposition: Define the set

$$\mathcal{H} := \left\{ \mathbf{x} \in \mathbb{R}^d \setminus \{\mathbf{0}_d\} \mid \|\mathbf{x}\| \leq h(\mathbf{x}/\|\mathbf{x}\|) \right\} \cup \{\mathbf{0}_d\},$$

where $\mathbf{0}_d := (0, \dots, 0)$. Then \mathcal{H} is star-shaped, compact, and satisfies $\mathcal{H} \subseteq [-1, 1]^d$, with unit-level set $\partial\mathcal{H} = \{\mathbf{w}h(\mathbf{w}) : \mathbf{w} \in \mathcal{S}^{d-1}\}$.

- To ensure $\partial\mathcal{H}$ also satisfies the **blue property**, we perform a numerical rescaling during the estimation procedure. This provides a new construction for valid (rescaled) gauge functions:

$$\tilde{g}(\mathbf{w}) := 1 / \left\| h(\kappa^{-1}(\mathbf{w})) \left(\frac{\kappa^{-1}(\mathbf{w})_1}{b_1(\kappa^{-1}(\mathbf{w})_1)}, \dots, \frac{\kappa^{-1}(\mathbf{w})_d}{b_d(\kappa^{-1}(\mathbf{w})_d)} \right) \right\|,$$

where $\kappa(\cdot)$ is a closed form bijective mapping and $b_i(\cdot), i = 1, \dots, d$, are scaling coefficients.

- We extend [5] and assume that, for any $\mathbf{w} \in \mathcal{S}^{d-1} \setminus \mathcal{A}$, where $\mathcal{A} := \bigcup_{i=1}^d \{\mathbf{w} \in \mathcal{S}^{d-1} : w_i = 0\}$ is the intersection of \mathcal{S}^{d-1} with each axis,

$$\Pr\left(\min_{i=1, \dots, d} \{X_i/w_i\} > u\right) \sim L(e^u; \mathbf{w}) \exp(-\Lambda(\mathbf{w})u), \quad u \rightarrow \infty, \quad (1)$$

where $\Lambda(\mathbf{w})$ denotes the (extended) angular dependence function (ADF).

- To link Λ to $\partial\mathcal{G}$, we have $\Lambda(\mathbf{w}) = \|\mathbf{w}\|_\infty \times \tilde{\tau}_{\mathbf{w}}^{-1}$, where $\tilde{\tau}_{\mathbf{w}} = \max\{\tau \in [0, 1] : \tau\tilde{\mathcal{R}}_{\mathbf{w}} \cap \partial\mathcal{G} \neq \emptyset\}$ and $\tilde{\mathcal{R}}_{\mathbf{w}} := \bigotimes_{i=1, \dots, d} \mathcal{U}_{w_i}$, with $\mathcal{U}_{w_i} := [w_i/\|\mathbf{w}\|_\infty, \infty]$ for $w_i > 0$ and $[-\infty, w_i/\|\mathbf{w}\|_\infty]$ for $w_i < 0$.

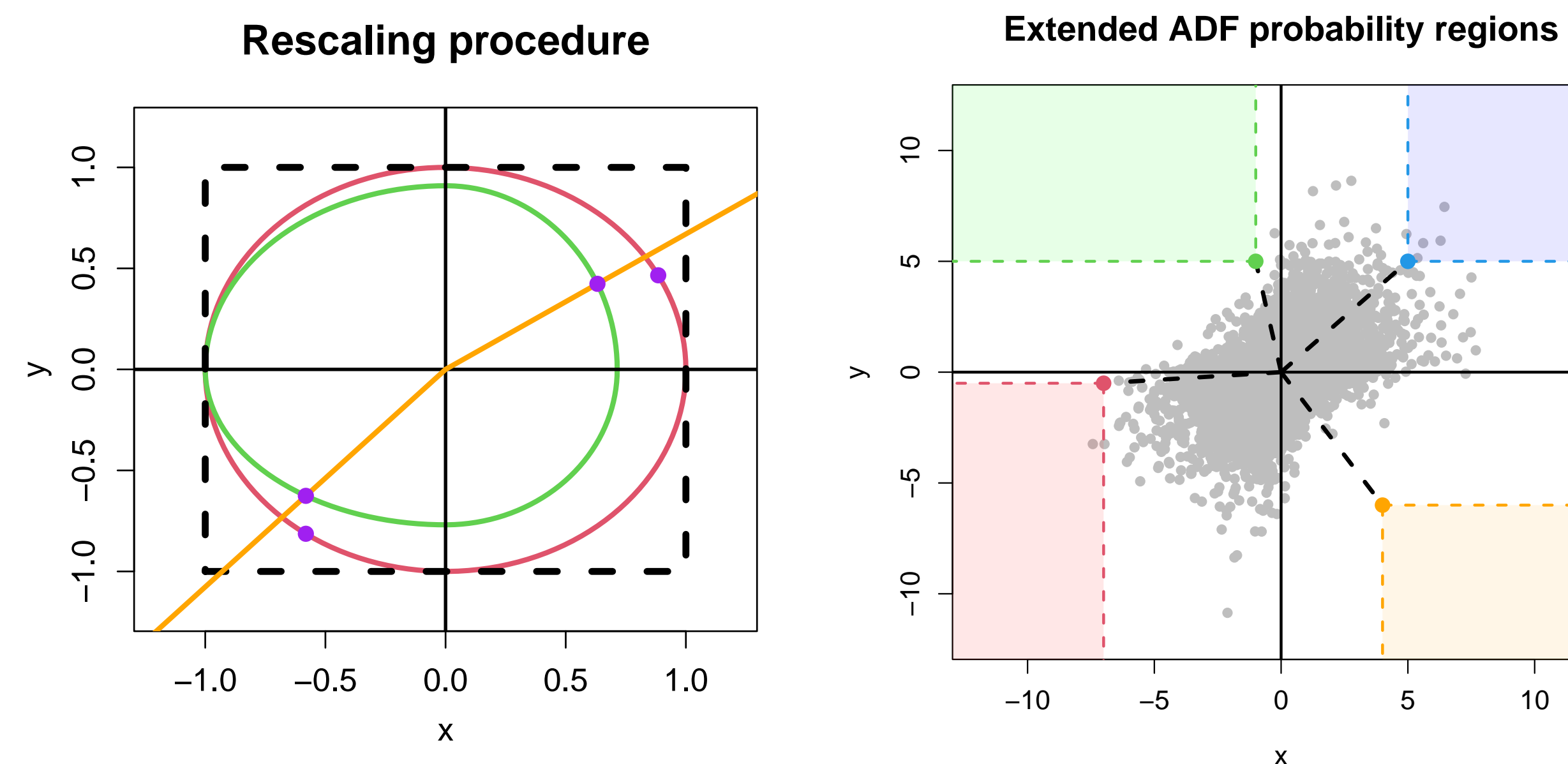


Figure 2. Left: rescaling. Right: example probability regions that can be evaluated using model (1).

Modelling and inference with neural networks

- For conditional radial distribution $R \mid (\mathbf{W} = \mathbf{w})$, we follow [6] and assume that

$$R \mid (\mathbf{W} = \mathbf{w}, R > r_\tau(\mathbf{w})) \sim \text{truncGamma}(\alpha, \tilde{g}(\mathbf{w})),$$

where $\alpha > 0$ and $r_\tau(\mathbf{w}) > 0$ satisfies $\Pr\{R \leq r_\tau(\mathbf{w}) \mid \mathbf{W} = \mathbf{w}\} = \tau$ for $\tau \in (0, 1)$ close to one.

- Then $\tilde{g}(\mathbf{w})$ is the rate parameter for the gamma distribution on $(R \mid \mathbf{W} = \mathbf{w}, R > r_\tau(\mathbf{w}))$ and can be found via standard maximum likelihood techniques.

- We model $r_\tau(\mathbf{w})$ and $1/h(\mathbf{w})$ as multi-layer perceptrons which take the angles \mathbf{w} as their input. These are designed so that $r_\tau(\mathbf{w}) > 0$ and $1/h(\mathbf{w}) \geq \|\mathbf{w}\|_\infty$ for all $\mathbf{w} \in \mathcal{S}^{d-1}$. The rescaling of $h(\cdot)$ is then performed to get the (rescaled) gauge function $\tilde{g}(\mathbf{w})$, $\mathbf{w} \in \mathcal{S}^{d-1}$.

- We estimate the conditional quantile $r_\tau(\mathbf{w})$ using standard quantile regression techniques. Inference for both this, and $\tilde{g}(\mathbf{w})$, is performed using the **R** interface to the deep learning library **Keras**.

Application to metocean extremes

- NORA10 hindcast dataset [Norwegian ReAnalysis 10km, 7];
- Gridded, 3-hourly fields at 10 km resolution;
- September 1957 – December 2009;
- Wind speed (ws; m/s), significant wave height (hs; m), mean sea level pressure (mslp; hPa);
- Uncertainty through stationary bootstrap.

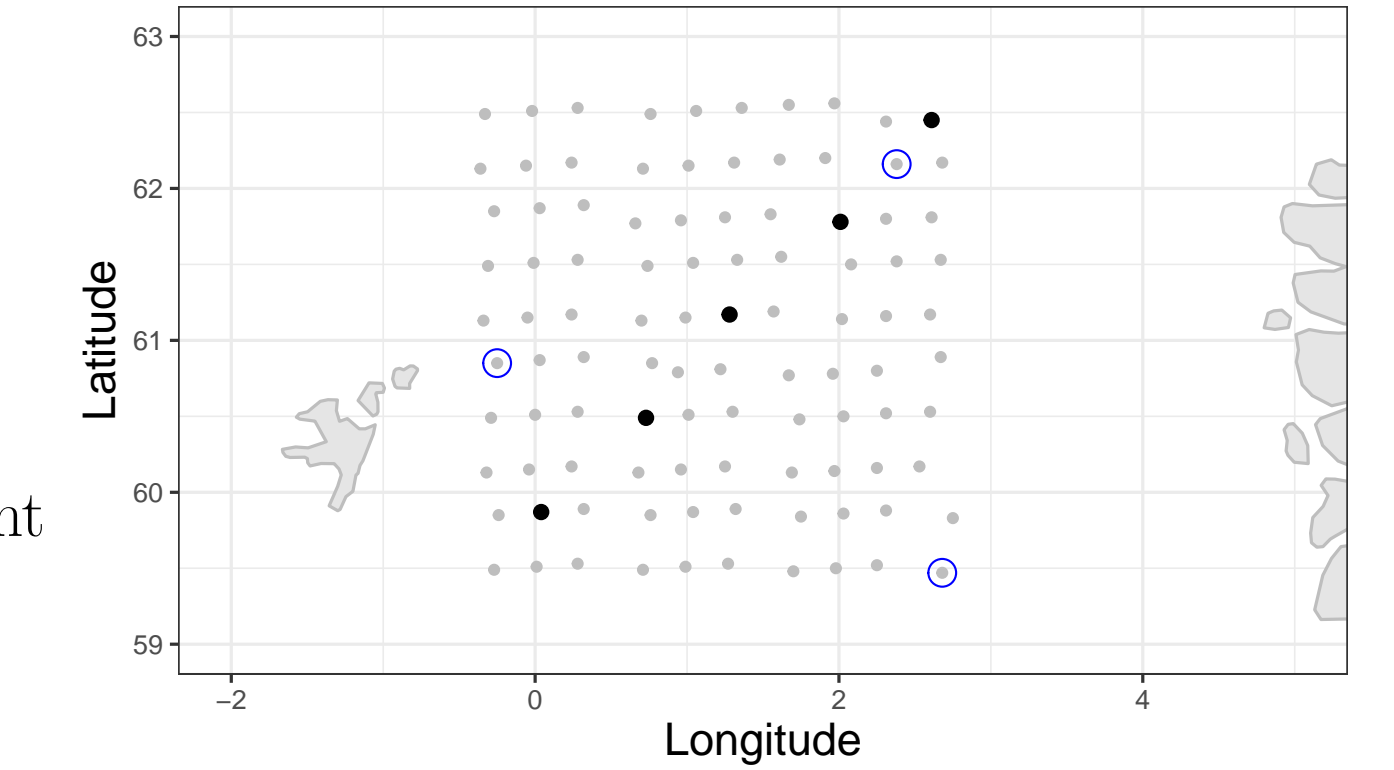


Figure 3. Study locations.

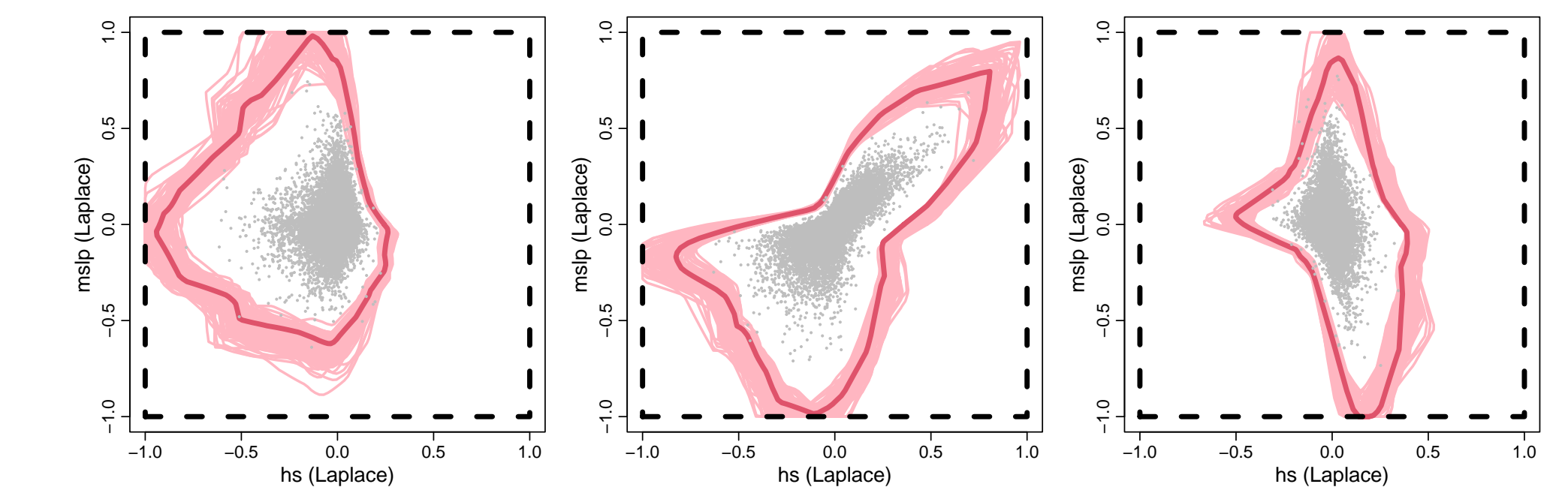


Figure 4. Estimated unit-level set slices for north-east location.

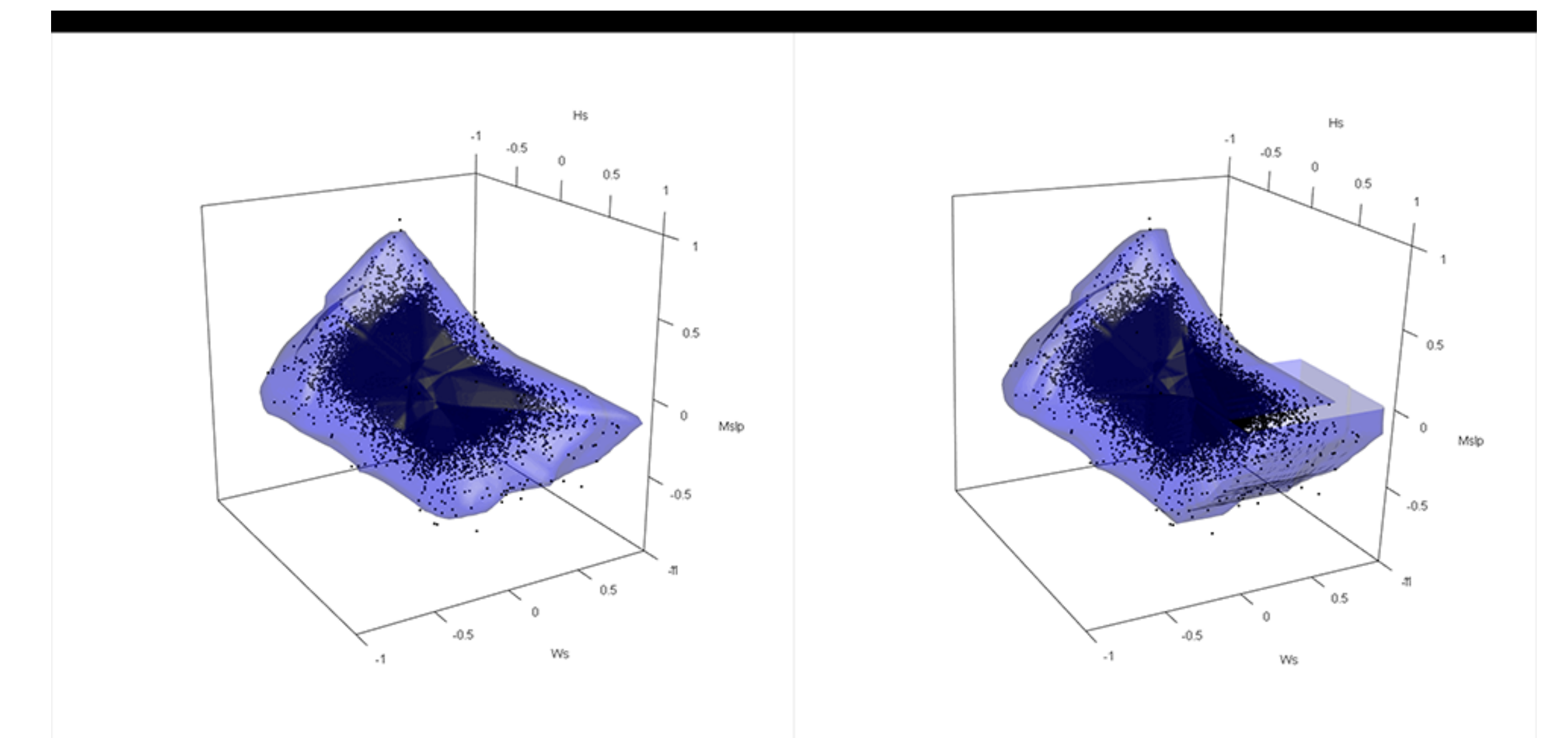


Figure 5. Estimated limit set (left) and $\Lambda(\cdot)$ (right) for north-east location.

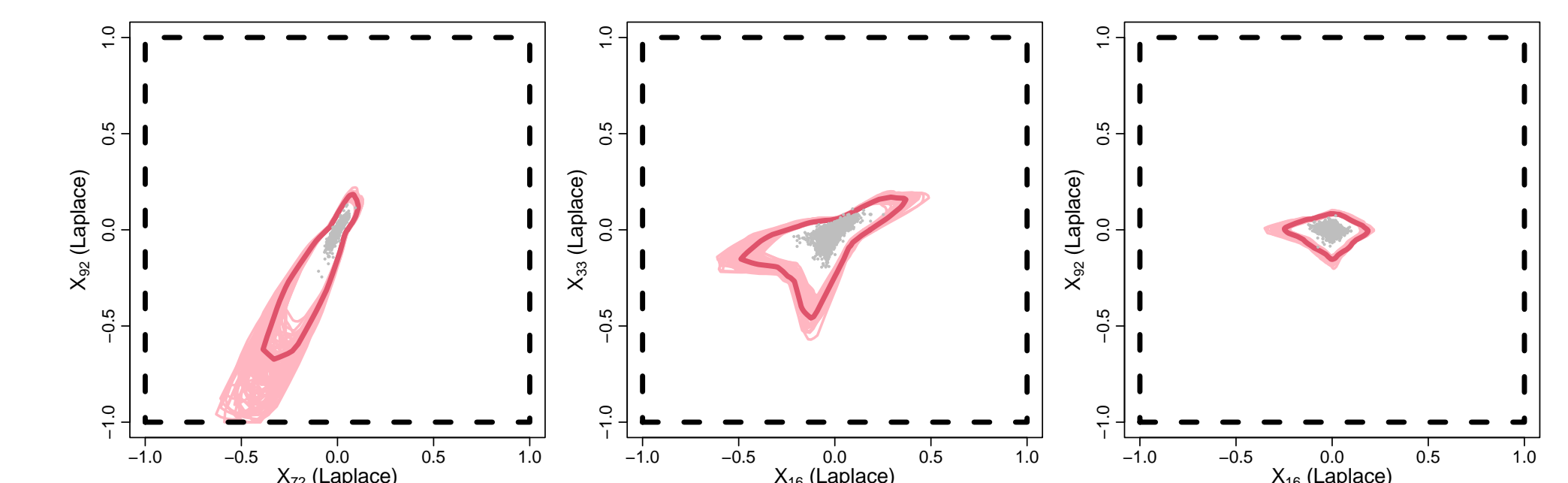


Figure 6. Estimated unit-level set slices for $d = 5$ transect (black locations).

References

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