

Introduction

- **Motivation:** existing literature on causality in multivariate extreme value theory tends to overlook the interaction between the upper and lower tails of random variables.
- **Contribution:** 1) we propose a new method to detect causal relationships for both tails of time series. 2) we contribute a high-frequency dataset from China's derivatives market.

Regular variation in a transformed space

- Regular variation $\mathbf{X} \in RV_+^p(\alpha)$ is often assumed in multivariate extreme statistics; \mathbf{X} has tail decay according to a power-law with index α and an angular measure $H_{\mathbf{X}}(\cdot)$ on the unit sphere $\mathbb{S}_{p-1}^+ = \{\mathbf{x} \in \mathbb{R}_+^p : \|\mathbf{x}\|_2 = 1\}$.
- With $t(x) = \log(\exp(x) + 1)$, the space $\mathbb{X}^p = \{t(\mathbf{x}) : \mathbf{x} \in \mathbb{R}^p\}$ is an inner product space (Cooley & Thibaud, 2019) with transformed linear operations (Table 1).

Operation	Definition
addition	$x_1 \oplus x_2 = t(t^{-1}(x_1) + t^{-1}(x_2))$
scalar multiply	$a \circ x = t(at^{-1}(x))$
inner product	$\langle x_1, x_2 \rangle = \sum_{i=1}^p t^{-1}(x_{1i})t^{-1}(x_{2i})$

Figure 1: Transformed operations

Regular variation is preserved in \mathbb{X}^p

- Tail Pairwise Dependence Matrix (TPDM) summarizes extremal dependence: $\Sigma_{\mathbf{X}} = \{\sigma_{ij}\}$, where $\sigma_{ij} = \int_{\mathbb{S}_{p-1}^+} w_i w_j dH_{\mathbf{X}}(\mathbf{w})$.
- Partial Tail Correlation Coefficient (PTCC):

$$\gamma_{ij|\mathbf{X}_{-ij}} = \frac{\langle X_i \ominus \widehat{X}_i, X_j \ominus \widehat{X}_j \rangle}{\|X_i \ominus \widehat{X}_i\| \|X_j \ominus \widehat{X}_j\|}; \mathbf{X}_{-ij} \text{ is } \mathbf{X} \text{ without } (X_i, X_j) \text{ and } \widehat{X}_i \text{ is the optimal transformed predictor for } X_i \text{ given } \mathbf{X}_{-ij}.$$

We can test for partial tail uncorrelatedness ($\gamma_{ij|\mathbf{X}_{-ij}} = 0$) denoted as $X_i \perp_p X_j | \mathbf{X}_{-ij}$, following (Lee & Cooley, 2022).

Transformed linear causal model

Let $X_i^u = T(X_i^{original})$, $X_i^d = T(-X_i^{original}) \in RV_+(2)$.

For $\mathbf{X}^{original} \in \mathbb{R}^p$, construct $\mathbf{X} = [X_1^u, X_1^d, \dots]' \in RV_+^{2p}(2)$.

Inference for one tail

Structured Causal Model (SCM) represented by a Directed Acyclic Graph (DAG) needs 1) a graph G ; 2) a causal equation system: $X_i^u := \eta_i^u \oplus \bigoplus_{X_j^e \in Pa(X_i^u)} X_j^e \circ \beta_{j \rightarrow i}^u$. The parameter space for this model is $\Theta_1 = (C, B)$, where C stores β and $B = \text{diag}(\sigma_{\eta_1}, \dots)$.

Theorem 1 $\Sigma_{\mathbf{X}^u} = (I - C)^{-1}B[(1 - C)^{-1}]'$ (system's TPDM)

We propose learning the SCM by PCMC1+ (Runge, 2020) and PTCC

Phase 1: Skeleton discovery - iteratively test $X_i \perp_p X_j | \mathbf{X}_{-ij}$

Phase 2: Edge orientation, which gives the causal direction

Application: Danube dataset

Danube data contains daily river discharges (n=428) from 31 stations. Here we estimate SCM and do comparison with: (Lee & Cooley, 2022), (Engelke & Hitz, 2020), and (Gong et al., 2022)

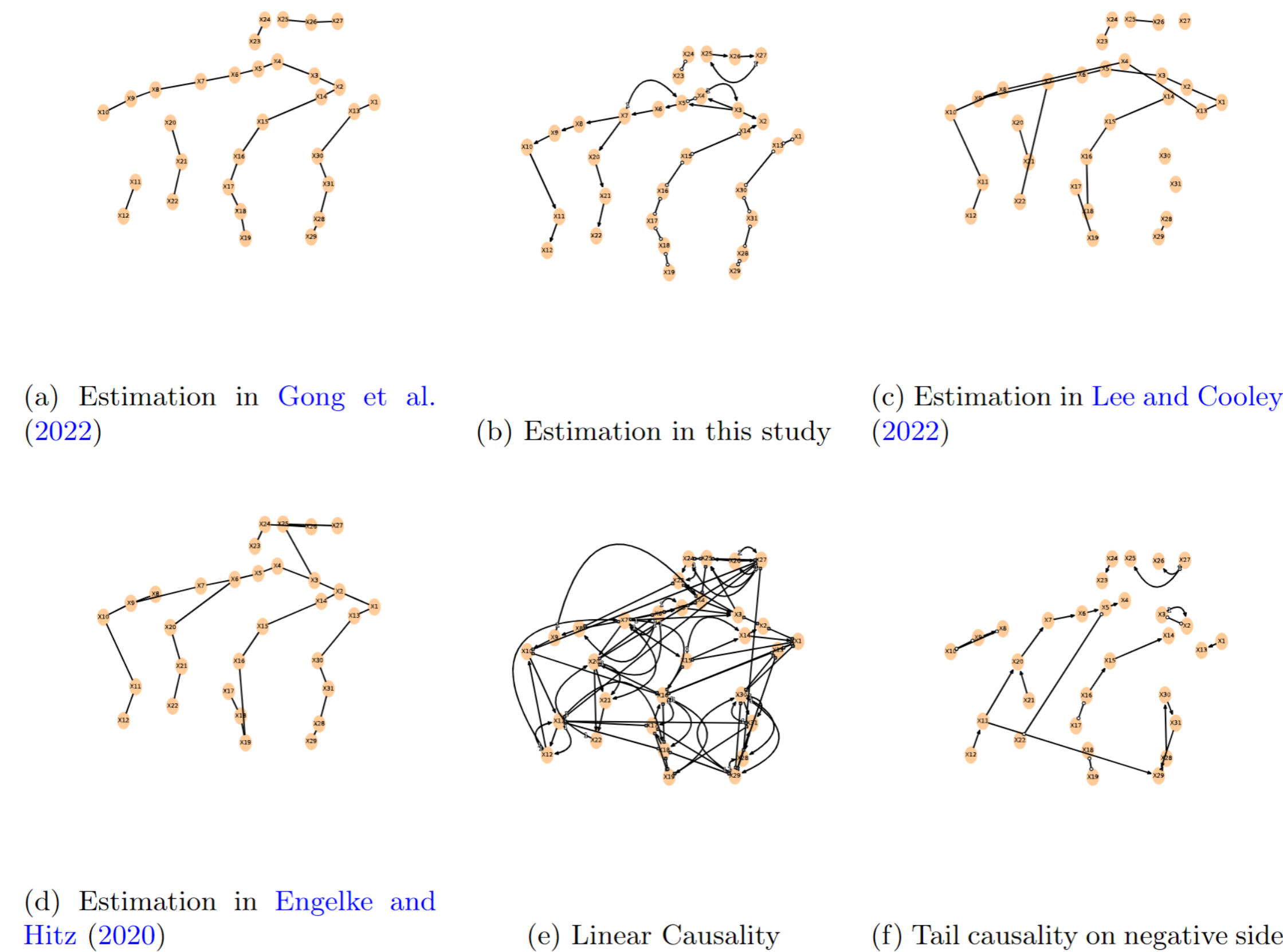


Figure 2: Comparison for Danube river discharge dataset; our model can show direction and temporal dependence.

Application: high-frequency China's derivatives market

New Dataset collection

- It is a thriving market (turnover is 87 trillion ¥ CNY in 2021)
- Include orderbook (2 records/sec); 80 million records a day;
- Potential research topics: data mining, risk-contagions, market efficiency and microstructure, high-frequency trading.

code	name	exchange	code	name	exchange
rb	Rebar	SHFE	b	Soybean Type.2	DCE
m	Soybean Meal	CFEEX	j	Coke	DCE
a	Soybean Type.1	DCE	jm	Coking Coal	DCE
al	Aluminum	SHFE	cu	Copper	SHFE
CY	Cotton Yarn	CZCE	y	Soybean Oil	DCE
i	Iron Ore	CZCE	zn	Zinc	SHFE

Figure 3: Derivative products for China's derivatives market

Estimation of causal structure

- 21 products, 4 sectors (color metal, chemicals,...).
- Results show one-step (5 seconds) forward SCM (Figure 4)
- Most are symmetric, suggesting market's efficiency in tails

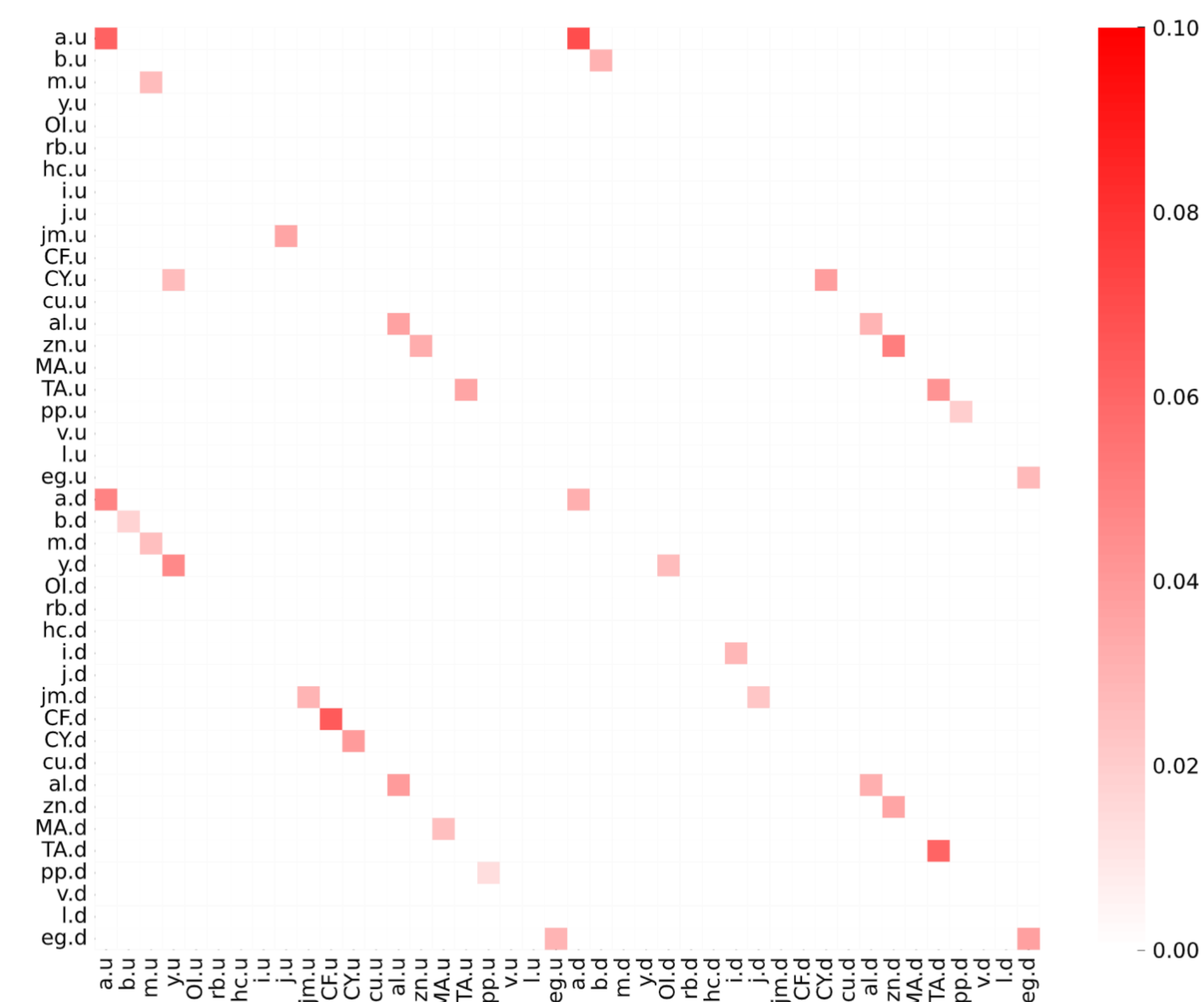


Figure 4: One-step forward causality for the China's derivatives return; quantified by PTCC.

Transformed linear causal model for both tails

- Since the orthogonality constraint $\sigma_{(X_i^u, X_i^d)} = 0$ exists:
- $X_i^u := I_i[\eta_i^u \oplus \bigoplus_{X_j^e \in Pa(X_i^u)} X_j^e \circ \beta_{j \rightarrow i}^u]$;
- $X_i^d := (1 - I_i)[\eta_i^d \oplus \bigoplus_{X_j^e \in Pa(X_i^d)} X_j^e \circ \beta_{j \rightarrow i}^d]$
- Swither $I_i = \mathbf{1}(X_i^{ori} > 0)$ are binary R.Vs (Illustrated in Figure 5).

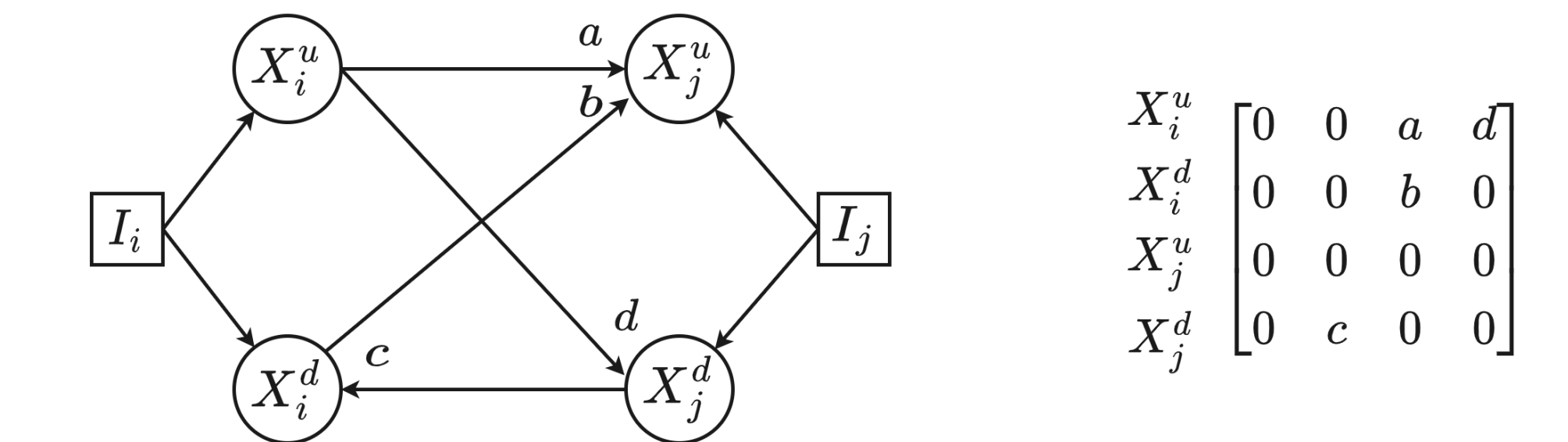


Figure 5: Left: one example of a causal graph for generalized transformed linear causality. Right: Causal matrix C for those 4 variables.

Conclusion

- We propose a transformed causal model which can nicely model the causal structure for both tails of time series data.
- Most of the lead-lag dependence relationship are symmetric in China's derivatives market.
- Interaction between both tails should receive more attention.
- We build a high-frequency dataset for China's derivatives market.

References

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