# A deep geometric approach to modelling multivariate extremes

#### Jordan Richards<sup>1</sup> Callum J. R. Murphy-Barltrop<sup>2,3</sup> Reetam Majumder<sup>4</sup>

<sup>1</sup>School of Mathematics, University of Edinburgh

<sup>2</sup>Technische Universität Dresden, Institut Für Mathematische Stochastik, Dresden <sup>3</sup>Center for Scalable Data Analytics and Artificial Intelligence (ScaDS.AI), Dresden/Leipzig

<sup>4</sup>Southeast Climate Adaptation Science Center, North Carolina State University

- コント 4 日 > ト 4 日 > ト 4 日 > - シックク

## Multivariate extremes

- Quantifying extremal dependence is key for risk analysis
- Many classical approaches require restrictive assumptions on joint tail decay, e.g., regular variation, max-stability
- Focus on particular extremes only



Multiple Variables

・ロト ・日 ・ ・ 日 ・ ・ 日 ・

Sac

## Geometric extremes

- Recent works have shown that deterministic limit sets provide a useful tool for studying extremal dependence (Nolde and Wadsworth, 2022; Mackay and Jonathan, 2023; Papastathopoulos et al., 2024).
- The "geometric approach" does not require restrictive parametric assumptions about joint tail decay.





ロト ( 目 ) ( 三 ) ( 三 ) ( ○ ) (

- All theory is applicable to *d*-dimensional random vectors  $X \in \mathbb{R}^d$  on standard margins, with density function  $f(\cdot)$ .
- Given *n* independent realisations of **X**, consider scaled sample cloud

$$C_n := \{\mathbf{X}_i/r_n; i = 1,\ldots,n\},\$$

- コント 4 日 > ト 4 日 > ト 4 日 > - シックク

as  $n \to \infty$ , where  $r_n$  is a suitably chosen normalising sequence. • For exponential (Laplace) margins,  $r_n = \log(n)$  ( $r_n = \log(n/2)$ ).

# Limit sets

#### Gaussian copula $\rho = 0.5$ with Laplace margins

<□ > < @ > < E > < E > E の < @</p>

Suppose

$$-\log f(t\mathbf{x}) \sim tg(\mathbf{x}), \ t \to \infty, \ \mathbf{x} \in \mathbb{R}^d,$$

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

where  $g(\cdot)$  is a continuous function on  $\mathbb{R}^d$ .

- $g(\mathbf{x})$  is termed the gauge function.
- $g(\mathbf{x})$  is 1-homogeneous, i.e.,  $g(c\mathbf{x}) = cg(\mathbf{x})$  for any c > 0.

• As  $n \to \infty$ ,  $C_n$  converges in probability onto the set

$$\mathcal{G} = \{\mathbf{x} \in \mathbb{R}^d : g(\mathbf{x}) \leq 1\} \subseteq [-1, 1]^d.$$

- G is star-shaped and compact (closed and bounded).
- For some centre  $\boldsymbol{o}$ , we have that the lines segment  $\{\boldsymbol{o} + t\boldsymbol{x} : t \in [0,1]\} \subset \mathcal{G}$  for all  $\boldsymbol{x} \in \mathcal{G}$ .
- Componentwise  $\max \mathcal{G} = (1, \dots, 1)^T$  and  $\min \mathcal{G} = (-1, \dots, -1)^T$ .

- コント 4 日 > ト 4 日 > ト 4 日 > - シックク

• See Nolde and Wadsworth (2022) for further details.

# Limit sets

Consider the unit level (boundary) set given by

$$\partial \mathcal{G} = \{ \mathbf{x} : g(\mathbf{x}) = 1 \} \subseteq [-1, 1]^d$$



х

▲ロト ▲ □ ト ▲ 三 ト ▲ 三 ト ● ● ● ● ●

#### Limit sets

 There exist links between ∂G and several existing approaches for multivariate extremes: Ledford and Tawn (1996), Heffernan and Tawn (2004), Wadsworth and Tawn (2013) and Simpson et al. (2020).

Once we have  $\partial \mathcal{G}$ , we get the rest for free.

 Whereas the above models focus on specific parts of the distribution, knowing ∂G gives you the complete picture of extremal dependence. • For example, the approach of Ledford and Tawn (1996) when **X**<sub>E</sub> has exponential margins:

$$\Pr\left(\min_{i=1,\ldots,d}\{X_{E,i}\}>u\right)\sim L(e^u)\exp(-u/\eta),$$

\*ロ \* \* @ \* \* ミ \* ミ \* ・ ミ \* の < @

as  $u \to \infty$ , with  $L(\cdot)$  slowly varying and  $\eta \in (0, 1]$ .

η quantifies the form of extremal dependence, with asymptotic dependence in X<sub>E</sub> implying η = 1.

# Ledford and Tawn

• We have that

$$\eta = \min\left\{s \in (0,1]: [s,\infty]^d \cap \partial \mathcal{G} = \emptyset
ight\}.$$

# Ledford and Tawn

#### Scaled sample cloud



х

▲□▶ ▲□▶ ▲ 三▶ ▲ 三▶ ● 三 ● ○ ○ ○

# Limit sets

#### Scaled sample cloud



х

#### Limit sets

*∂Ĝ* gives risk metrics; return level sets (Papastathopoulos et al., 2024), return curves (Murphy-Barltrop et al., 2024) and joint tail probabilities (Wadsworth and Campbell, 2024).

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

- Knowing  $\partial \mathcal{G}$  is invaluable for inference.
- How do we estimate ∂G?

Many recent approaches have been proposed for estimating  $\partial \mathcal{G}$ .

- Simpson and Tawn (2022) used generalised additive models to approximate  $\partial \mathcal{G}$  via scaled radii sets.
- Wadsworth and Campbell (2024) propose truncated parametric copula models for estimating  $\partial \mathcal{G}$ .
- Majumder et al. (2023) proposed a semi-parametric approach with Bézier polynomials.

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

• Papastathopoulos et al. (2024) provided a Bayesian inference approach using latent Gaussian variables.

Many more approaches (likely) to follow.

Most require parametric or semi-parametric assumptions about the gauge function  $g(\cdot)$ :

- Limited to generally low dimension  $d \leq 3$ ;
- Inflexible;
- $\bullet$  Do not necessarily guarantee that estimates of  $\partial {\cal G}$  satisfy the red and blue properties

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

# Talk outline

- $\bullet$  Build valid estimators of  $\partial \mathcal{G}$  using flexible neural networks
- Provide a means to estimate probabilities using an extension of the Wadsworth and Tawn (2013) ADF

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

Application to metocean extremes

# Radial-angular decomposition

Define the radial and angular components

$$R := \|\mathbf{X}\|_2, \quad \mathbf{W} := \frac{\mathbf{X}}{R},$$

\*ロ \* \* @ \* \* ミ \* ミ \* ・ ミ \* の < @

so  $\mathbf{X} = R\mathbf{W}$ .

• Observe  $\|\boldsymbol{W}\|_2 = 1$ . Therefore,  $\boldsymbol{W}$  will always exist on the unit (d-1)-sphere  $\mathcal{S}^{d-1} := \{ \mathbf{x} \in \mathbb{R}^d : \|\mathbf{x}\|_2 = 1 \}.$ 

## Data

#### Laplace margins



・ロト・西ト・山下・山下・山下・山下

## Polar coordinates

#### Angular-radial decomposition



 • By star-shaped property of  $\mathcal{G}$ , we can re-characterise the unit-level set:

$$\partial \mathcal{G} = \{ r \boldsymbol{w} : r > 0, \boldsymbol{w} \in \mathcal{S}^{d-1}, g(r \boldsymbol{w}) = 1 \}.$$

• By homogeneity of  $g(\cdot)$ , we must have  $r = 1/g(\boldsymbol{w})$ , implying

$$\partial \mathcal{G} = \{ \boldsymbol{w}/g(\boldsymbol{w}) : \boldsymbol{w} \in \mathcal{S}^{d-1} \}.$$

▲□▶ ▲□▶ ▲臣▶ ▲臣▶ 三臣 - のへで

• To evaluate  $\partial \mathcal{G}$ , we just need to evaluate  $g(\cdot)$  over the sphere  $\mathcal{S}^{d-1}$ .

# Limit set

#### Scaled sample cloud



х

・ロト < 団ト < 三ト < 三ト < 三 ・ < ロト</li>

# Limit set: radially transformed

#### , 0.8 0.4 0.0 5 0 2 3 4 6 $\cos^{-1}(w_1)$ Gaussian copula, rho = 0.5 Red: $1/g(\mathbf{w})$ . Black: upper-bound for $1/g(\mathbf{w})$ .

#### Angular-radial decomposition

# Bound on g

#### Proposition

For all  $\boldsymbol{w} \in S^{d-1}$ , we have that  $g(\cdot)$  satisfies

 $g(\boldsymbol{w}) \geq ||\boldsymbol{w}||_{\infty},$ 

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

where  $||\mathbf{x}||_{\infty} := \max\{|x_1|, \dots, |x_d|\}$  denotes the infinity norm.

## Bound on g : valid estimators

Suppose we have any  $h(\cdot) : S^{d-1} \mapsto \mathbb{R}_+$  satisfies  $1/h(\boldsymbol{w}) \ge ||\boldsymbol{w}||_{\infty}$  for all  $\boldsymbol{w} \in S^{d-1}$ . Then:

Proposition

Define the set

$$\mathcal{H} := \bigg\{ \boldsymbol{x} \in \mathbb{R}^d \setminus \{\boldsymbol{0}_d\} \ \bigg| \ ||\boldsymbol{x}|| \leq h(\boldsymbol{x}/||\boldsymbol{x}||) \bigg\} \bigcup \bigg\{ \boldsymbol{0}_d \bigg\},$$

- コント 4 日 > ト 4 日 > ト 4 日 > - シックク

where  $\mathbf{0}_d := (0, \dots, 0)$ . Then  $\mathcal{H}$  is star-shaped, compact, and satisfies  $\mathcal{H} \subseteq [-1, 1]^d$ .

## Bound on g : valid estimators

The corresponding unit-level set  $\partial \mathcal{H}$  is

$$\partial \mathcal{H} = \left\{ \boldsymbol{w} h(\boldsymbol{w}) : \boldsymbol{w} \in \mathcal{S}^{d-1} 
ight\}.$$

- Whilst  $\partial \mathcal{H} \subseteq [-1, 1]^d$ , it might not satisfy coordinate-wise max and min property of  $\partial \mathcal{H}$ , i.e., max  $\partial \mathcal{H} = \mathbf{1}_d$  and min  $\partial \mathcal{H} = -\mathbf{1}_d$
- Can be easily obtained using a straightforward rescaling
- Implication:starting with a general radial function  $h(\cdot)$  satisfying  $h(\mathbf{w}) \ge ||\mathbf{w}||_{\infty}$ , we can construct valid unit level sets

# Rescaling

For each 
$$i = 1, ..., d$$
, we define  
 $b_i(w_i) := \mathbb{1}(w_i \ge 0)b_i^U - \mathbb{1}(w_i < 0)b_i^L > 0$ , where  
 $b_i^U := \max\left\{w_i h(\boldsymbol{w}) \mid \boldsymbol{w} \in S^{d-1}\right\} > 0$ , and  
 $b_i^L := \min\left\{w_i h(\boldsymbol{w}) \mid \boldsymbol{w} \in S^{d-1}\right\} < 0$ .

Using these scaling functions, we define the rescaled set

$$\widetilde{\partial \mathcal{H}} := \left\{h(\boldsymbol{w})\left(\frac{w_1}{b_1(w_1)}, \ldots, \frac{w_d}{b_d(w_d)}\right) \, \middle| \, \boldsymbol{w} \in \mathcal{S}^{d-1}\right\}.$$

<ロト < 部 ト < 注 ト < 注 ト 三 三 のへで</p>



#### Proposition

The rescaled set  $\partial \mathcal{H}$  is in one-to-one correspondence with  $\partial \mathcal{H}$ , satisfies  $\partial \mathcal{H} \subset [-1,1]^d$ , and has componentwise maxima and minima  $\mathbf{1}_d$  and  $-\mathbf{1}_d$ , respectively.

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

# Rescaling

#### **Rescaling procedure**



х

◆□ > ◆□ > ◆ 臣 > ◆ 臣 > ◆ 臣 = ∽ へ ⊙

## Rescaling

This provides a new construction for valid (rescaled) gauge functions

$$ilde{g}(oldsymbol{w}) := 1 \Big/ \left\| h(\kappa^{-1}(oldsymbol{w})) \left( rac{\kappa^{-1}(oldsymbol{w})_1}{b_1(\kappa^{-1}(oldsymbol{w})_1)}, \dots, rac{\kappa^{-1}(oldsymbol{w})_d}{b_d(\kappa^{-1}(oldsymbol{w})_d)} 
ight) 
ight\|$$

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

where  $\kappa(\cdot)$  is a bjiective mapping.

- In practice, we model  $h(\cdot)$  using a neural network;
- Compute scaling factors *b<sub>i</sub>* numerically;
- By uniform sampling (many) points on  $S^{d-1}$ .

With estimates of  $\tilde{g}(\cdot)$ , we can get tail probability estimates.

# Ledford and Tawn

• Recall L+T (96):

$$\Pr\left(\min_{i=1,\ldots,d}\{X_{E,i}\}>u\right)\sim L(e^u)\exp(-u/\eta),$$

<ロト < 部 ト < 注 ト < 注 ト 三 三 のへで</p>

as  $u \to \infty$ , with  $L(\cdot)$  slowly varying and  $\eta \in (0, 1]$ .

## Angular dependence function (ADF)

Wadsworth and Tawn (2013) generalise L+T (96) with the angular dependence function  $\lambda(\mathbf{w})$ .

• For angle  $\boldsymbol{w} = (w_1, \dots, w_d)^T \in \mathcal{S}^{d-1}_+$ :

$$\Pr\left(\min_{i=1,\ldots,d}\{X_{E,i}/w_i\}>u\right)\sim L(e^u;\boldsymbol{w})\exp(-\lambda(\boldsymbol{w})u)$$

as  $u \to \infty$ , with  $\mathcal{S}^{d-1}_+ := \{ \boldsymbol{x} \in \mathbb{R}^d_+ : ||\boldsymbol{x}|| = 1 \}.$ 

•  $\lambda(\boldsymbol{w})$  quantifies extremal dependence along different rays  $\boldsymbol{w}$ . • With  $\eta^{-1} = \{\sqrt{d}\lambda(d^{-1/2}, \dots, d^{-1/2})\}.$ 

- Assume W + T (2013) model holds for any reflection  $\boldsymbol{cX}$ , where  $\boldsymbol{c} \in \{-1,1\}^d$  ( $\boldsymbol{X}$  on Laplace margins).
- For any  $\boldsymbol{w} \in S^{d-1} \setminus A$ , where  $A := \bigcup_{i=1}^{d} \{ \boldsymbol{w} \in S^{d-1} : w_i = 0 \}$  is the intersection of  $S^{d-1}$  with each axis:

$$\Pr\left(\min_{i=1,\ldots,d}\{X_i/w_i\}>u\right)\sim L(e^u;\boldsymbol{w})\exp(-\Lambda(\boldsymbol{w})u),\quad u\to\infty,$$

where  $\Lambda(\boldsymbol{w})$  denotes the (extended) ADF.

- Quantifies strength of extremal dependence along any *w* ∈ S<sup>d-1</sup> \ A (not S<sup>d-1</sup><sub>+</sub>) as moved from exponential to Laplace margins.
- See, also, Mackay and Jonathan (2023).

Unit-level set  $\partial \mathcal{G}$ , alongside  $\{\boldsymbol{w}/\Lambda(\boldsymbol{w}): \boldsymbol{w} \in \mathcal{S}^{d-1} \setminus \mathcal{A}\}.$ 



<ロト < 団 ト < 三 ト < 三 ト < 三 の < ()</p>

We prove the following links between  $\Lambda$  and  $g/\partial \mathcal{G}$ :

• 
$$g(\boldsymbol{w}) \ge \Lambda(\boldsymbol{w}) \ge ||\boldsymbol{w}||_{\infty}$$
;  
•  $\Lambda(\boldsymbol{w}) = ||\boldsymbol{w}||_{\infty} \times \tilde{\mathfrak{r}}_{\boldsymbol{w}}^{-1}$ , where  $\tilde{\mathfrak{r}}_{\boldsymbol{w}} = \max\{\mathfrak{r} \in [0,1] : \mathfrak{r}\tilde{\mathcal{R}}_{\boldsymbol{w}} \cap \partial \mathcal{G} \neq \emptyset\}$   
and  $\tilde{\mathcal{R}}_{\boldsymbol{w}} := \bigotimes_{i=1,...,d} \mathcal{U}_{w_i}$ , with  $\mathcal{U}_{w_i} := [w_i/||\boldsymbol{w}||_{\infty}, \infty]$  for  $w_i > 0$  and  
 $[-\infty, w_i/||\boldsymbol{w}||_{\infty}]$  for  $w_i < 0$ .

Can be used to estimate probabilities of the form

$$\Pr(\operatorname{sgn}(x_i)X_i > \operatorname{sgn}(x_i)x_i, i = 1, \ldots, d).$$

▲□▶▲□▶▲□▶▲□▶ □ ● ● ●

## Inference

To estimate  $\tilde{g}(\cdot)$ :

- Consider  $R \mid (\mathbf{W} = \mathbf{w})$ .
- Following Wadsworth and Campbell (2024), assume

$$R \mid (\boldsymbol{W} = \boldsymbol{w}, R > r_{\tau}(\boldsymbol{w})) \sim \operatorname{truncGamma}(\alpha, \tilde{g}(\boldsymbol{w})),$$

where  $\alpha > 0$  and  $r_{\tau}(\boldsymbol{w}) > 0$  satisfies  $\Pr\{R \leq r_{\tau}(\boldsymbol{w}) \mid \boldsymbol{W} = \boldsymbol{w}\} = \tau$  for  $\tau \in (0, 1)$  close to one.

• Then  $\tilde{g}(\boldsymbol{w})$  is the rate parameter for the gamma distribution on  $(R \mid \boldsymbol{W} = \boldsymbol{w}, R > r_{\tau}(\boldsymbol{w})).$ 

Our inference framework has two steps:

- Estimation of the threshold function  $r_{\tau}(\boldsymbol{w})$ ,  $\boldsymbol{w} \in \mathcal{S}^{d-1}$ .
- Estimation of the (rescaled) gauge function  $\widetilde{g}(\boldsymbol{w})$ ,  $\boldsymbol{w} \in \mathcal{S}^{d-1}$ .

We perform both estimation steps using multilayer perceptrons, say  $m(\cdot)$ , which take input **w**.

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

# Neural net setup

- The threshold r<sub>τ</sub>(·) > 0 is represented as exp{m(w)} and estimated first using quantile regression.
- We model  $h(\mathbf{w}) = \operatorname{ReLU}(m(\mathbf{w})) + ||\mathbf{w}||_{\infty}$  to ensure  $h(\mathbf{w}) \geq ||\mathbf{w}||_{\infty}$ , and then rescaled to get  $\tilde{g}(\cdot)^1$ .
- Inference is performed using Keras for R.

<sup>1</sup>*ReLU*(
$$\mathbf{x}$$
) = max{ $x$ , 0}.

# Case study

- NORA10 hindcast dataset (NOrwegian ReAnalysis 10km, Reistad et al., 2011)
- Gridded product, 3-hourly wave fields at 10 km res.
- covering the Norwegian Sea, the North Sea, and the Barents Sea
- September 1957 December 2009
- wind speed (ws) measured in m/s, significant wave height (hs) measured in m, and mean sea level pressure (mslp) measured in hPa;

- コント 4 日 > ト 4 日 > ト 4 日 > - シックク

# Case study





Figure: Pairwise ws, hs, and mslp at one location.



Figure: Pairwise distributions of hs across four locations in transect.

# Diagnostics - d = 3



Figure: QQplots for fit of  $\tilde{g}(\cdot)$  and  $\Lambda(\cdot)$  for location 85.

◆□ > ◆□ > ◆豆 > ◆豆 > ・豆 - つへ⊙

## Results - d = 3

Figure: Limit set (left) and  $\Lambda(\cdot)$  (right) for location 46.

#### Results - d = 3



Figure: Unit level set slices for location 46.

◆□▶ ◆圖▶ ◆注▶ ◆注▶ ─ 注一

990

## Results - d = 5



Figure: Unit level set slices for transect.

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > <

Ξ

990

# Conclusions

- "Deep Learning of Multivariate Extremes via a Geometric Representation" is available on arxiv (2406.19936)
- Further theoretical results for limit sets (with Laplace margins)
- Includes simulation study and practical advice for validating model fits

<ロト 4 目 ト 4 三 ト 4 三 ト 9 0 0 0</p>

R code available soon!

Fin.



Scan for full details of my research.

<ロト < 部 ト < 注 ト < 注 ト 三 三 のへで</p>

### References I

- Heffernan, J. E. and Tawn, J. A. (2004). A conditional approach for multivariate extreme values. *Journal of the Royal Statistical Society. Series B: Statistical Methodology*, 66:497–546.
- Ledford, A. W. and Tawn, J. A. (1996). Statistics for near independence in multivariate extreme values. *Biometrika*, 83:169–187.
- Mackay, E. and Jonathan, P. (2023). Modelling multivariate extremes through angular-radial decomposition of the density function. *arXiv*, 2310.12711.
- Majumder, R., Shaby, B. A., Reich, B. J., and Cooley, D. (2023). Semiparametric estimation of the shape of the limiting multivariate point cloud. arXiv, 2306.13257.
- Murphy-Barltrop, C. J. R., Wadsworth, J. L., and Eastoe, E. F. (2024). Improving estimation for asymptotically independent bivariate extremes via global estimators for the angular dependence function. *arXiv*, 2303.13237.

### References II

- Nolde, N. and Wadsworth, J. L. (2022). Linking representations for multivariate extremes via a limit set. Advances in Applied Probability, 54:688–717.
- Papastathopoulos, I., de Monte, L., Campbell, R., and Rue, H. (2024). Statistical inference for radially-stable generalized pareto distributions and return level-sets in geometric extremes. arXiv, 2310.06130.
- Reistad, M., Øyvind Breivik, Haakenstad, H., Aarnes, O. J., Furevik, B. R., and Bidlot, J.-R. (2011). A high-resolution hindcast of wind and waves for the north sea, the norwegian sea, and the barents sea. *Journal* of Geophysical Research, 116:C05019.
- Simpson, E. S. and Tawn, J. A. (2022). Estimating the limiting shape of bivariate scaled sample clouds: with additional benefits of self-consistent inference for existing extremal dependence properties. arXiv, 2207.02626.

- Simpson, E. S., Wadsworth, J. L., and Tawn, J. A. (2020). Determining the dependence structure of multivariate extremes. *Biometrika*, 107:513–532.
- Wadsworth, J. L. and Campbell, R. (2024). Statistical inference for multivariate extremes via a geometric approach. *Journal of the Royal Statistical Society Series B: Statistical Methodology*, 2208.14951.

Wadsworth, J. L. and Tawn, J. A. (2013). A new representation for multivariate tail probabilities. *Bernoulli*, 19:2689–2714.